

1 **Title:** Optimal sampling methods for modelling the occupancy of Arctic Grayling (*Thymallus arcticus*) in  
2 the Canadian Barrenlands

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15 **Number of tables and figures:** 1 table, 5 figures

16 **Number of references:** 30

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24 **Abstract**

25 In occupancy models, imperfect detectability of animals is usually corrected for by using temporally-  
26 repeated surveys to estimate probability of detection. Substituting spatial replicates for temporal  
27 replicates could be an advantageous sampling strategy in remote Arctic regions, but may lead to serious  
28 violations of model assumptions. Using a case study of site occupancy of adfluvial young-of-year Arctic  
29 Grayling in Barrenland tundra streams, we assessed reliability and efficiency of alternative sampling  
30 strategies; i) randomly distributed vs sequential adjacent spatial replicates; ii) visual vs electrofishing  
31 surveys; and, iii) spatial vs temporal replicates. Sequential, adjacent spatial replicates produced spatially  
32 auto-correlated data. Autocorrelation was relieved using randomly distributed spatial replicates, but  
33 using these randomly distributed spatial replicates introduced significant error into estimates of the  
34 probability of occupancy in streams. Models designed for spatially-autocorrelated data could minimize  
35 this bias. Visual and electrofishing surveys produced comparable probabilities of detection. Spatially-  
36 replicated surveys performed better than temporal replicates. The easiest and relatively most cost-  
37 effective sampling methods performed as well as, or better than, the more established, expensive, and  
38 logistically difficult alternatives for occupancy estimation.

39

40 **Key-words:** Arctic, detection probability, correlated detections, electrofishing, occupancy model,  
41 multiscale model, salmonid, spatial replication, streamside visual surveys

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47 **Introduction**

48           Freshwater ecosystems in the Arctic are experiencing rapid change in climate, and increasing  
49 pressure from ever-growing industrial development. The impacts of anthropogenic stressors on  
50 hydrology, water temperature, primary productivity, food web structure, and fish life history are  
51 expected to be far-reaching, but have been poorly quantified in these remote and under-studied  
52 ecosystems (Prowse et al. 2012, Reist et al. 2006a, Wrona et al. 2006). Comprehensive, standardized  
53 datasets are needed for larger-scale integration of data (Reist et al. 2006b), but studies to date on  
54 northern fishes have used a multitude of methods and data analysis tools that preclude synthesis on  
55 large spatial or temporal scales. Reliable monitoring programs can be costly in terms of both financial  
56 and personnel resources; thus, the development of a cost-effective data collection framework for  
57 sensitive northern fish populations is essential to their conservation.

58           In the Barrenlands region, adfluvial populations of Arctic Grayling (*Thymallus arcticus*, Pallas),  
59 like other migratory fishes, are sensitive to fragmentation or alterations of the habitats they utilize (Reist  
60 et al. 2006b). Young-of-year Arctic Grayling hatch and rear for several months in clear, cool, gravel or  
61 rock-bottomed streams (Scott and Crossman 1973) before migrating to overwintering sites in lakes  
62 (Jones and Tonn 2004). The Barrenlands landscape is a priority research area for many northern  
63 stakeholders including industries, regulators, and Indigenous groups working to mitigate effects of mine-  
64 and/or climate-related stream dewatering on populations of adfluvial Arctic Grayling. Despite this,  
65 habitat use by young-of-year, adfluvial Arctic Grayling in the Barrenlands has only been investigated in a  
66 handful of streams (e.g., Jones and Tonn 2004).

67           Occupancy, defined as the proportion of area, patches, or sample units that are occupied (i.e.,  
68 species presence) by a given species, is a natural state variable that can be used in studies of species  
69 distribution and range. Occupancy models are a means of deriving information regarding the ecological

70 niche of a species (e.g., Hutchinson 1957), as each species has a unique set of requirements that must  
71 be provided by habitats used. Identification of key habitat variables that species respond to can be used  
72 to develop habitat models that predict patch and landscape-level occupancy (e.g., see Verner *et al.*  
73 1986; Scott *et al.* 2002). In remote northern environments, it may be particularly advantageous to apply  
74 an occupancy modeling framework (MacKenzie *et al.* 2002) to monitoring programs of landscapes that  
75 are too large and logistically difficult to survey extensively. Time and effort spent sampling a site can be  
76 reduced by focusing sampling efforts on collection of presence-absence data (instead of abundance  
77 data) in a manner that permits inference to the entire area of interest, allowing greater spatial and  
78 temporal coverage of a species' distribution across the landscape (Royle and Nichols 2003). Occupancy  
79 modeling also explicitly addresses issues of imperfect detection (i.e., false absence) (MacKenzie *et al.*  
80 2002). Failing to account for false absences can introduce significant error into species distribution  
81 models (Gu and Swihart 2004).

82         The standard method for estimating detection probabilities in occupancy studies involves  
83 surveying a site multiple times over a defined 'season' (MacKenzie *et al.* 2002). Temporally-replicated  
84 surveys can be expensive and logistically difficult to implement in remote areas, and resources invested  
85 in visiting the same site multiple times within a given timeframe and budget limits spatial coverage of  
86 survey efforts. This may be especially problematic in surveys of Arctic fishes because repeat visits must  
87 be made within the relatively short ice-free season. Alternatively, the replicate surveys may take the  
88 form of randomly-selected spatial replicates within the sample site. Spatial replication is relatively less  
89 costly than multiple site visits, but occupancy of each replicate must be independent of the other  
90 replicates within the site (Hines *et al.* 2010), and there must be uniform availability of the species for  
91 detection in all spatial sub-units of an occupied site (Kendall and White 2009). These assumptions may  
92 be violated when, for example, fish exhibit non-random spatial distributions due to schooling behaviour

93 or when they are subject to downstream displacement in stream ecosystems during high water flow  
94 events. As a result of the potential for violation of assumptions, some authors have cautioned against  
95 the use of spatial replicates instead of temporal replicates (Kendall and White 2009), yet the actual  
96 amount of bias induced by use of spatial replicates in an occupancy study has rarely been quantified  
97 using real data. Occupancy models that include a first-order Markovian occupancy process (Gillespie  
98 1992), in which the probability of occupancy in a spatial replicate  $j$  depends on whether the species was  
99 present or absent from the previous spatial replicate  $j-1$ , have been developed (Hines *et al.* 2010) to  
100 handle issues where replicate spatial surveys suffer from a sequential form of spatial autocorrelation,  
101 such as may be present when replicates are constrained to linear landscape features like streams.

102         The probability of detecting a species can also be influenced by the sampling method used  
103 (Nichols *et al.* 2008). Backpack electrofishing and visual counts from streambanks are two commonly  
104 used fish detection techniques. The relative efficiency of these two methods in producing abundance  
105 estimates is well-characterized (Bozek and Rahel 1991), however, their efficiency in collecting presence-  
106 absence data for occupancy studies has not been addressed. Both techniques can suffer from bias  
107 resulting from fish size and behaviour, and can only be used in relatively shallow, (<1 m) clear water  
108 (Ensign *et al.* 2002). Electrofishing techniques require less observer standardization, but the electrical  
109 current can harm fish (Dwyer and White 1997, Reynolds 1996). Streamside visual surveys are less likely  
110 to result in altered behaviour or harm to fish (Brewer and Ellersieck 2011), but the identification of  
111 cryptically-coloured fish from the stream bank requires a greater level of skill (Bozek and Rahel 1991).  
112 Electrofishing gear is typically heavy (10-15 kg), expensive, and can be difficult to use in remote, rugged  
113 terrain, whereas streamside visual surveys do not require the operator to enter the water (when  
114 streams are narrow enough) and do not require any specialized equipment (Bozek and Rahel 1991).

115           Occupancy models were developed using field observations of adfluvial young-of-year Arctic  
116 Grayling near a diamond mine development in the Northwest Territories, Canada (DeBeers' Gahcho  
117 Kué). The objectives of this study were to quantify bias in occupancy models that results from  
118 alternative sampling methods, and specifically compare: 1) models of data derived from surveys of  
119 sequential, adjacent spatial replicates to models of data derived from randomly-selected spatial  
120 replicates; 2) relative detection probabilities of two commonly used observational techniques for  
121 freshwater fishes (backpack electrofishing vs. streamside visual); 3) relative efficiency of using only  
122 spatial vs. only temporal replicates to estimate site occupancy; and, 4) using the best models, examine  
123 Arctic Grayling young of-year occupancy patterns in streams as they related to habitat characteristics  
124 and industrial activities.

125

## 126 **Methods**

### 127 **Case study area**

128           The Kennady Lake drainage system is located approximately 280 km north northeast of  
129 Yellowknife, Northwest Territories, Canada (63°26'15 N, 109°11'51 W) (Fig. 1) within the sub-Arctic  
130 Tundra Shield ecozone. Situated north of the treeline, it is part of a vast area commonly referred to as  
131 the Barrenlands region; a semi-arid sub-arctic landscape with low levels of precipitation (between 200-  
132 300 mm annually - over half of which falls as snow; (Environment Canada 1991)). The development of a  
133 new open-pit diamond mine (Gahcho Kué), required draining a section of Kennady Lake. Prior to  
134 development, Kennady Lake provided overwintering habitat for an adfluvial population of Arctic  
135 Grayling, as well as several other fish species. The adfluvial Arctic Grayling in this system likely will  
136 continue to use the undrained portion of Kennady Lake as overwintering habitat, in addition to several  
137 other downstream chain lakes. The study area encompassed Barrenland streams ranging 90-800 m in

138 length, with each end connected to lakes, over approximately 100 km<sup>2</sup>. The study area includes streams  
139 within the Kennady Lake drainage basin, the Kirk lake drainage basin and the Walmsley Lake drainage  
140 basin in the Northwest Territories of Canada.

141

## 142 **Field survey methods**

143           Prior to the start of the dewatering of Kennady Lake in 2014, baseline data of occupancy of  
144 Arctic Grayling young-of-year in streams were collected. Sixty-seven stream segments (segments=spatial  
145 replicates of streams that were each 30 m in length) in nine streams (KLM system; Fig. 1) downstream of  
146 Kennady Lake were surveyed four times each during the summer of 2014. In summer 2015, after lake  
147 dewatering had begun, 105 segments in 20 streams were surveyed up to three times in three areas: i)  
148 streams immediately downstream of Kennady Lake, now affected by dewatering (the original KLM  
149 systems, n=9 streams); ii) streams further downstream of Kennady Lake but less likely to have been  
150 affected by dewatering (the P system, n=5 streams); and, iii) streams in a reference watershed not  
151 affected by dewatering, downstream of Walmsley Lake (the W system; n=6 streams) (Fig. 1).

152           To quantify the bias introduced to spatially-replicated stream occupancy models by organisms  
153 exhibiting a lack of independence in their spatial distribution (project objective 1), sequential adjacent  
154 stream segments were surveyed in 2014. The entire length of each stream in the KLM system was  
155 surveyed in 30-m segments, of all streams (which served as the spatial replication within the stream)  
156 (see Fig. S1a in Supporting Information). Results were compared to those generated by surveying a  
157 random selection of segments in 2015; up to six stratified, randomly-selected 30-m segments were  
158 surveyed in each stream (Fig. S1b) instead of entire streams. It was necessary to collect these data sets  
159 in separate years for them to be considered independent, where *a posteriori* resampling of the data  
160 would be equivalent to non-parametric bootstrapping (Efron and Tibshirani 1994), which tends to return

161 the same point estimate as the original data (Kendall and White 2009). To compare relative efficiency of  
162 using spatial vs. temporal replicates to estimate stream occupancy (objective 3), we collected both  
163 spatially-replicated and temporally-replicated survey data in each survey year (Fig. S1a and b).

164 To quantify the detection efficiency of two common fish detection techniques (project objective  
165 2), we used both techniques to independently detect fish in all surveys. Field surveys were conducted by  
166 moving upstream from the furthest downstream end of each stream. Polarized sunglasses were worn  
167 during surveys to reduce glare from the water surface. Streamside visual surveys were conducted from  
168 streambanks. Observations were conducted by two observers simultaneously from opposite banks of  
169 the stream. These surveys were combined to a single observation of presence or absence of Arctic  
170 Grayling young-of-year was recorded for each stream segment (i.e. each spatial replicate). Quantitative  
171 estimates of variables that may affect the probability of detection were recorded, including cloud cover  
172 and surface visibility (glare and turbulence; see Table S1). Starting again at the furthest downstream end  
173 of each stream, single-pass electrofishing surveys were conducted moving upstream using a Smith-Root  
174 LR-20B backpack electrofisher with a 6-inch anode ring (Voltage – 990 V; Duty Cycle – 50%; Frequency –  
175 35, 0.20 A output). The backpack operator and netter moved upstream together using a zig-zag pattern  
176 to shock fish, sampling micro-habitats proportionally. Low specific conductivity of stream water (10-15  
177  $\mu\text{S cm}^{-1}$ ) limited the effective range of the electrofisher to approximately 2 m. Presence or absence of  
178 Arctic Grayling young-of year was recorded for each 30-m stream segment.

179 Habitat variables hypothesized to affect occupancy or detection of Arctic Grayling young-of-year  
180 were collected from each stream. Covariates that were expected to affect detection of fish were  
181 assessed at the scale of the individual survey (see Table S1), whereas covariates expected to affect the  
182 occupancy of fish in streams were measured at each stream segment (except discharge and distance to  
183 overwintering habitat), and averaged to produce a single value representative of the entire stream.



184 Proportion of stream margins with floodplain wetlands (defined as the presence of shallow, standing  
185 surface water over hydric soils, adjacent to the main stream channel, (Tiner 1999)), proportion of stream  
186 margins with undercut banks (defined as a stable bank which overhangs a stream (Dohner et al. 1997),  
187 and percent cover of vegetation types (emergent, submerged and good overhanging vegetation (Nielson  
188 and Johnson 1983)) were estimated visually. Stream width (tape measure), depth, and velocity (Hach  
189 FH950 handheld flow meter mounted on a wading rod) were also quantified. Distance to overwintering  
190 habitat was assessed as the minimum number of lake crossings required for Arctic Grayling young-of-  
191 year to reach lakes with overwintering habitat (<4 m in depth which included Kennady Lake, Lake M4,  
192 Lake 410, Kirk Lake and Walmsley Lake). Discharge was assessed using the United States Geological  
193 Survey mid-section method (adapted from (Buchanan and Somers 1969) at a single fixed location for  
194 each stream. All covariates were standardized to z-scores prior to analysis and checked for excessive  
195 collinearity. Those found to be highly correlated (correlations  $\geq 0.50$ ) were not included together in a  
196 single model, but considered only in competing models to prevent overestimation of probability of  
197 occupancy or detection.

198

### 199 **Statistical analysis**

200 Following MacKenzie et al. (2002), models of probability of Arctic Grayling young-of-year  
201 occupancy in streams (herein the term “probability of occupancy” will always refer to occupancy of  
202 Arctic Grayling young-of-year in streams, unless otherwise specified) were assessed using the occupancy  
203 modeling estimation and information theoretic approach. To estimate the relative utility of using  
204 sequential spatial replicates vs randomly-selected spatial replicates (objective 1), and the relative utility  
205 of using visual surveys vs. electrofishing surveys (objective 2), we modelled the probability of occupancy  
206 of Arctic Grayling young-of-year with hierarchal models that utilized all data from spatially- and

207 temporally-replicated surveys (data configurations shown in Fig. S1 a and b) for each observational  
208 technique (visual and electrofishing) and for each year (2014 and 2015). Four hierarchical data sets were  
209 created and used in modelling: 2014-visual hierarchical, 2014-electrofishing hierarchical, 2015-visual  
210 hierarchical and, 2015-electrofishing hierarchical. Each of these four data sets was modeled by testing  
211 the relative fit of three *a priori* candidate model structures and evaluated using the adjusted Akaike  
212 information criterion ( $AIC_c$ ; using the number of stream segments as the sample size). The difference in  
213  $AIC_c$  values was used to provide a relative weight-of-evidence for each candidate model structure ( $w_i$ ) for  
214 each data set. All modeling was performed using the program PRESENCE v10.7 (Hines 2006).

215         The first candidate hierarchical spatial-temporal model structure was a simple “multi-season”-  
216 style model,  $[\psi(\cdot), \gamma(\cdot), \epsilon(\cdot), \rho(\cdot)]$ , (referred to as Candidate Model 1: Open Occupancy, see Supplemental  
217 Data S1 in Supporting Information for more details on each of the candidate models and explanations of  
218 variables). These models are typically used when surveys are repeated annually (or “seasons”), but in  
219 the present study we treated each survey period within each of 2014 and 2015 as a “season”. Spatial  
220 replication within each season was used to assess probability of detection, and results were used to  
221 provide guidance on the most appropriate timing for surveys if only one spatially-replicated survey were  
222 to be conducted in each year. Two additional candidate models were used to approximate a possible  
223 lack of independence in occupancy of segments within streams, which would test whether spatial  
224 heterogeneity of fish in the streams existed, and was not explained by habitat covariates, and if the  
225 randomly-selected spatial replicates survey style alleviated spatial dependency. A multi-scale occupancy  
226 model,  $[\psi(\cdot), \theta(\text{segment}), \rho(\cdot)]$ , (referred to as Candidate Model 2: Clustered Spatial Correlation) was  
227 used to approximate nested spatial scales in the sampling design; stream segments ( $\theta$ ) were nested  
228 within streams ( $\psi$ ), and streams were nested within survey period. A multi-season Markovian occupancy  
229 model structure,  $[\psi(\cdot), \theta_0(\cdot), \theta_1(\cdot), \gamma(\cdot), \epsilon(\cdot), \rho(\cdot), \theta_0 * \pi(=0)]$ , (referred to as Candidate Model 3: Sequential

230 Spatial Correlation with Open Occupancy) was used to test downstream spatial autocorrelation in the  
231 occupancy of replicate segments within streams. In this model, a first order Markovian spatial process is  
232 used; the probability of occupancy of Arctic Grayling young-of year in a stream is decomposed into three  
233 components - occupancy,  $\psi$ , and two availability variables given absence or presence in the adjacent  
234 stream segment,  $\theta_0$  or  $\theta_1$ , respectively (Hines 2010).

235           The relative rankings of the three candidate models described above were used to address  
236 objective 1. If spatial heterogeneity existed in the occupancy of fish in streams, then candidate model 2  
237 (clustered occupancy) or 3 (sequentially clustered occupancy) should rank highest by AIC for the 2014-  
238 visual hierarchical and the 2014-electrofishing hierarchical data sets in which sequential spatial surveys  
239 were used (Hines et al. 2010). If spatial heterogeneity exists in the occupancy of fish in the study  
240 streams, and the use of randomly-selected spatial replicated surveys, as were used in 2015, ameliorated  
241 the effect of this spatial heterogeneity, then candidate model 1 (open occupancy-no clustering in  
242 occupancy) should rank as the highest model by AIC for the 2015-visual hierarchical and 2015-  
243 electrofishing hierarchical data sets, suggesting that randomly-selected spatial segments may be a useful  
244 survey design for Arctic stream fish.

245           Using the best of the three candidate hierarchical models as selected by AIC<sub>c</sub> ranking, a  
246 benchmark hierarchical model was produced for each of the 4 data sets using a sequential model-  
247 building strategy to account for non-random (i.e., resulting from biological or measurement covariates)  
248 variation in probability of occupancy or detection. The probability of occupancy was modelled as a  
249 function of stream-level biological covariates, and the probability of detection was modelled as a  
250 function of segment-level measurement parameters. First, a detection ( $p$ ) model was built using all  
251 subsets of covariates for the detection parameter (2014=2 detection covariates, 4 models; 2015=4  
252 detection covariates, 16 models), while holding all other model parameters constant. Occupancy models

253 were then constructed using all covariates singly (due to small sample sizes) on the large-scale  
254 occupancy parameter,  $\psi$  (2014=12 covariates, 48 models; 2015=8 covariates, 32 models), while holding  
255  $p$  at the most parsimonious model. Multi-model inference was achieved by averaging  $\beta$  parameter  
256 estimates and estimated probability of occupancy of streams ( $\psi$ ) of all models having  $\Delta AIC_c$  estimates  
257 within 2 of the top-ranked model (Richards 2005). Unconditional standard errors were estimated using  
258 the delta method (Falke et al. 2012). The importance of covariates was estimated based on the relative  
259 difference of model-averaged  $\beta$  estimates from zero (0=no importance). Beta coefficients for these  
260 benchmark hierarchical models are presented in Table S2 in the Supporting Information.

261 Objective 2 of this study was assessed by comparing the mean ( $\pm$  95% confidence intervals)  
262 probabilities of detection of Arctic Grayling young-of-year produced by the best 2014-visual hierarchical  
263 model vs. the best 2014-electrofishing hierarchical model, and by comparing probabilities of detection  
264 produced by the best 2015-visual hierarchical model vs. the best 2015-electrofishing hierarchical model.  
265 Because we cannot know the true probability of detecting Arctic Grayling young-of-year, we were only  
266 able to assess how similar the probabilities of detection for each observational method were to each  
267 other and how small the range of error was for each observational method. If the two observational  
268 methods produced similar probabilities of detection within the same year, then the prudent choice of  
269 the “best” observational method would be the one that is relatively less expensive in terms of effort and  
270 money, and produces the smallest amount of error in the estimates of probability of detection and  
271 occupancy.

272 We addressed objective 3 by comparing relative amount of bias in probability of occupancy  
273 produced from models applied to a simulated temporal-replicate-only data set (Fig. S1 c) vs a  
274 representative spatial-replicate-only data set (Fig. S1 d). Data from the 2015 survey campaign were  
275 used. Here, bias refers to differences in probability of occupancy relative to estimates produced by the

276 assumed best (benchmark) hierarchical spatial-temporal model. To simulate a temporally-replicated  
277 data set, presence/absence data from each spatial replicate within a stream were condensed to a single  
278 presence/absence data point that represented the entire stream for each of the three sampling periods  
279 in 2015. This data set thus consisted of presence/absence data for 20 streams visited up to three times  
280 in 2015. To represent a spatially-replicated data set, data from the second sampling period of 2015 were  
281 used. This data set consisted of presence/absence data for 20 streams, with up to six segments surveyed  
282 without temporal replication. Four datasets were thus produced from the 2015 survey data, 2015-  
283 visual-temporal only data (configuration c in Fig. S1), 2015-visual-spatial only data (configuration d in  
284 Fig. S1), 2015-electrofishing-spatial only data (configuration c) and, 2015-electrofishing-temporal only  
285 data (configuration d). Small sample size ( $n=9$  streams) precluded conducting the same analysis on data  
286 collected in 2014. Each of these data sets was modeled by evaluating the relative fit of two *a priori*  
287 candidate model structures using  $AIC_c$ , which were single season versions of the candidate models  
288 described in the previous model set. The two candidate models included a simple single-season model,  
289  $[\psi(\cdot), p(\cdot)]$ , and a single-season with correlated detections model,  $[\psi(\cdot), \theta_0(\cdot), \theta_1(\cdot), p(\cdot), \theta_0 * \pi(=0)]$ , and  
290 the sequential model-building strategy outlined earlier in the methods was used. Probabilities of  
291 occupancy produced by the best model of each of the 2015-visual-temporal only data (configuration c,  
292 Fig. S1) and 2015-visual-spatial only data (configuration d, Fig. S1) were compared to the probabilities of  
293 occupancy produced by the benchmark hierarchical model of the 2015-visual hierarchical data  
294 (configuration b, Fig. S1). This was accomplished by calculating the root mean square deviance (RMSD)  $\pm$   
295 95% confidence intervals. Again, we assumed that the benchmark hierarchical occupancy models  
296 produced the truest estimates of site occupancy. Similarly, probabilities of occupancy of streams  
297 produced by the best model of each of the 2015-electrofishing-temporal only data (configuration c, Fig.  
298 S1) and the 2015-electrofishing-spatial only data (configuration d, Fig. S1) were compared to the

299 probabilities of occupancy produced from the benchmark hierarchical model of the 2015-electrofishing  
300 hierarchical data set (configuration b, Fig. S1).

301 AIC assumes that the candidate model set contains at least one model that fits the data  
302 adequately; AIC is used to select the best model, but this is no assurance that the selected model is a  
303 good model, and substantial lack of fit can lead to inaccurate inferences (Anderson et al. 1994). Given  
304 the relative novelty and complexity of the models used in this study, robust methods available for  
305 testing the goodness-of-fit of the models have not yet been developed (pers. communication, D.  
306 MacKenzie). A qualitative testing procedure was used to indirectly assess the goodness of fit of the *a*  
307 *priori* candidate (or global) model for the hierarchical benchmark model selection for each of the 4  
308 hierarchical data sets (Cooch 2012). If the fit of the benchmark global models (which contain all possible  
309 parameters) is adequate, all subsets of these models are assumed to also fit the data because they  
310 originate from the global model (Burnham and Anderson 2002). The quasi-likelihood estimation  
311 parameter (QAICc, (Wedderburn 1974) is typically calculated as a correction for overdispersion based on  
312 the parametric bootstrapped goodness-of-fit chi-squared statistic ( $\hat{c}$ ). We arbitrarily set the  $\hat{c}$  to values  
313 of 1 (perfect fit) to 3 (overdispersed), in increments of 0.25, to see how this affected the relative ranking  
314 of candidate models. By adjusting  $\hat{c}$  to higher values, suggestive of a lack of fit of the models, the model  
315 selection becomes more conservative, which tends to favour models with less parameters. If  
316 overdispersion exists within the model set, the relative weightings and order of the candidate models  
317 change with small changes in  $\hat{c}$ , indicating a lack of fit of the *a priori* model structures, and indicating  
318 that the data may be too sparse for robust modelling. We found that the rankings of the *a priori*  
319 candidate sets did not change with changes in  $\hat{c}$ , lending some measure of confidence that the top-  
320 ranked models were a reasonable fit for the data (Cooch 2012).

321

322 **Results and discussion**

323 **Occupancy modelling using sequential, adjacent spatial replicates vs. randomly-selected spatial**  
324 **replicates**

325           The most supported model of the 2014 visual and electrofishing hierarchical data sets was  
326 Candidate Model 3: Sequential Spatial Correlation with Open Occupancy, indicating that sequential  
327 spatially replicated surveys of adjacent stream segments produced spatially auto-correlated data sets  
328 (Table 1). The presence of Arctic Grayling young-of-year in each stream segment was likely influenced by  
329 the presence of young-of-year in the upstream segment. In 2015, the selection of random segments  
330 resulted in segments being separated by an average distance 20 m (or approximately 0.7 segments,  
331 where one segment=30 m). Spatial auto-correlation was apparently relieved by the random spatial  
332 replicate selection process implemented in 2015, as the AIC analysis of the 2015 hierarchical data sets  
333 (both visual and electrofishing) indicated that the best supported model was Candidate Model 1: Open  
334 Occupancy. Thus, when adjacent spatial replicates were surveyed, the presence of fish in these  
335 replicates was not independent of the presence of fish in the upstream segment. However, when we  
336 instead surveyed only a subset of randomly-selected, non-adjacent spatial replicates the spatial  
337 dependence of the presence of fish in spatially replicated surveys was ameliorated. These findings  
338 suggest that either the area occupied by an interacting group of Arctic Grayling young-of-year, or the  
339 relative size of suitable summer rearing habitat patches used by groups of Arctic Grayling young-of-year  
340 in a stream, could be larger than 30 m (the size of the segments used as replicates), but smaller than 50  
341 m (the average distance between replicates in 2015 plus the size of the replicate). The home range of  
342 adfluvial populations of adult European Grayling (*Thymallus thymallus*, a sister species of the Arctic  
343 Grayling) in streams has been observed to be approximately 75-100 m, although daily movements  
344 ranged between 15-18 m (Nykänen et al. 2004). The home range of adfluvial populations of Arctic

345 Grayling appears to be less well characterized, but the typical size of cohesive groups of interacting  
346 Arctic Grayling young-of year in an Alaskan stream ranged between 4-52 m (Hughes and Reynolds 1994).

347

#### 348 **Probability of detecting fish with visual vs. electrofishing observational methods**

349 In both 2014 and 2015, probabilities of detecting fish using streamside visual surveys were  
350 nearly identical to those using electrofishing surveys (Fig. 2A and 2B). Probability of detection with  
351 electrofishing surveys was 3.3 % higher in 2014 and 3.4 % lower in 2015 than with visual surveys  
352 ( $p=0.0004$ ,  $n=4$  temporal replicates and  $p=0.0003$ ,  $n=3$  temporal replicates, respectively, paired t-tests).  
353 While these results are statistically significant, we believe that a 3-4% difference in detection probability  
354 is trivial, and that either survey method would produce similar quality of data. Overall, detection  
355 probability was higher but more variable in 2014, averaging  $54 \pm 5\%$ , compared to  $40 \pm 2\%$  in 2015.

356 In the surveys performed prior to the start of the dewatering of Kennady Lake (2014 surveys),  
357 there was improved probability of detection at water velocities above 10 cm/s, and the effect of water  
358 velocity on detection was nearly identical between the two sampling methods (Fig. 2C). This may reflect  
359 the somewhat poor swimming ability of fry at higher water velocities. Small Arctic Grayling young-of-  
360 year are poor swimmers and have previously been observed to prefer water velocities between 0-10 cm  
361  $s^{-1}$  (Jones and Tonn 2004). At water velocities above this preferred range, Arctic Grayling young-of-year  
362 may have been easier to detect because they were less able to swim quickly to a refugium in the higher  
363 water velocities.

364 There was no apparent effect of water velocity on probability of detecting fish after the start of  
365 dewatering in 2015. Average water velocity in stream segments was higher in 2015 ( $33 \text{ cm s}^{-1}$ ) than in  
366 2014 ( $8 \text{ cm s}^{-1}$ ), and was above the apparent threshold of  $10 \text{ cm s}^{-1}$  for maximum probability of detection  
367 in 2014 (Fig. 2C). The depth of stream segments was the only variable that appeared to affect probability



368 of fish detection in 2015 (Fig. 2D). Observers likely had greater difficulty in detecting fish in deeper  
369 waters; the magnitude of this effect was greater for streamside visual surveys than for electrofishing  
370 surveys. Stream segment depths in the KLM system were on average 10 cm deeper in 2015 (26-54 cm)  
371 than in 2014 (18-44 cm), where 10 cm total was observed as the optimum water depth for Arctic  
372 Grayling young-of-year in another Barrenlands stream system (Jones and Tonn 2004). There was likely  
373 much less habitat of suitable depth available in 2015, which may explain why depth affected probability  
374 of detection in 2015 but not in 2014. The increase in water depth of the KLM streams in 2015 was likely  
375 partially a result of mine operations; water from Kennady Lake was pumped across a berm into a lake  
376 that drains into stream K5 (Fig. 1). Natural hydrological variability could also have affected stream depth.  
377 Water depths in the KLM system in 2015 were within the range of water depths observed in the P and W  
378 systems (see Table S1), and summer precipitation was higher in 2015 (at 79.2 cm) than in 2014 (at 58.4  
379 cm) (Environment Canada 2016). Summer precipitation can strongly influence runoff and flooding into  
380 streams in the Barrenlands region (Marsh et al. 2008).

381         Streamside visual survey methods produced lower estimates ( $11 \pm 4\%$  lower) of the probability  
382 of occupancy than electrofishing methods. Although estimates of the probability of occupancy  
383 generated by models of streamside visual surveys were more variable than those generated by  
384 electrofishing surveys, the estimates of the probability of occupancy from streamside visual surveys  
385 were overall more similar to the naïve observations of fish presence in streams (Fig. 3A and B), and  
386 better able to distinguish sites where Arctic Grayling young-of-year appeared to be absent. Currently,  
387 electrofishing is regarded as the most effective monitoring technique of fish assemblages (Poos et al.  
388 2007), however, the present study suggests that this convention may not hold when the monitoring goal  
389 is landscape-scale presence-absence of fish, as opposed to estimates of abundance. While subtle  
390 differences in the two observational methods were apparent, we think that ultimately, the visual and

391 electrofishing surveys produced similar enough estimates of probabilities of both detection and  
392 occupancy that they could be considered as equivalent methods in terms of quality of data produced.  
393 However, the streamside survey method offers several logistical advantages. Streamside surveys are  
394 much less likely to disrupt or injure to fish, the cost of purchase and transport of gear is minimal,  
395 observers are not required to maneuver with heavy gear in the stream, and two observers can conduct  
396 independent streamside surveys, effectively doubling the data produced per unit of survey effort. In  
397 contrast, electrofishing surveys require two observers (an operator and a netter) to conduct a single  
398 survey.

399

#### 400 **Relative bias in occupancy models when using spatially replicated surveys vs. temporally replicated** 401 **surveys**

402 Estimates of probability of occupancy and detection produced from models of only spatially-replicated  
403 data better represented the benchmark hierarchical models (having open occupancy) than models using  
404 only the temporally-replicated data. Detection probabilities were comparable between the hierarchical  
405 and spatially-replicated data sets ( $p$  of  $\sim 0.50$ ), whereas the temporally-replicated data sets appeared to  
406 have much higher detection probabilities than the benchmark hierarchical models (Fig. 4A). The  
407 overestimation of detection probabilities in the temporally-replicated models likely resulted from  
408 combining the data from all spatial replicates into a hypothetical single survey; the probability of  
409 detection for the temporally-replicated model applies at the scale of the stream whereas the probability  
410 of detection for the spatially-replicated and benchmark hierarchical models apply at the scale of the  
411 spatial replicate; the 30-m segment. As such, we do not suggest that differences in probability of  
412 detection between spatially and temporally replicated models should be interpreted as one method  
413 producing better probabilities of detection over the other.

414 Both the single period of spatially-replicated surveys and the condensed temporally-replicated  
415 surveys produced positively biased and more variable probabilities of occupancy of streams than the  
416 benchmark hierarchical model. On average, models of the spatially-replicated streamside visual surveys  
417 overestimated the proportion of streams occupied by  $28.7 \pm 5\%$  compared to the hierarchical  
418 spatially/temporally-replicated model (Fig. 4B). Spatially-replicated electrofishing surveys overestimated  
419 probabilities of occupancy by  $32.1 \pm 9.4\%$  compared to hierarchical spatially/temporally-replicated  
420 electrofishing surveys (Fig. 4B). Temporally-replicated streamside visual and electrofishing surveys  
421 resulted in greater overestimations of the probability of occupancy ( $49.6 \pm 5.9\%$  and  $43.3 \pm 12.7\%$ ,  
422 respectively; Fig. 4B). Due to unequal sample sizes (spatial:  $n=6$ , and temporal:  $n=3$ ) these results do not  
423 necessarily disagree with previous occupancy studies, which report that spatial replication may not be a  
424 robust substitute for temporal replicates (Kendall and White 2009). Models of the spatially-replicated  
425 streamside visual survey data set were re-run using only 3 replicates, and the overestimation of the  
426 probability that streams are occupied that was produced by the equalized replication of spatial surveys  
427 increased from 28.7% to 40.3%; however, this is still a better estimate of the probability of occupancy  
428 than the temporally-replicated streamside visual surveys (at 49.6% overestimated probability of  
429 occupancy relative to the benchmark model).

430 Bias (compared to the benchmark hierarchical models) in the estimates of probability of  
431 occupancy was greater for streams in the system affected by the draining of the upstream lake (KLM  
432 system) compared to the control (P and W) streams (Fig. 4B). Spatially-replicated streamside visual  
433 surveys produced the most consistent (although still somewhat overestimated) estimates of probability  
434 of occupancy of streams in the KLM system and the control streams. All other combinations of survey  
435 method and replications failed to detect the probable decline in stream occupancy in the KLM system  
436 resulting from alteration of water flow in the area downstream of Kennedy Lake in 2015 (see presence

437 data in Fig 3). In any monitoring scenario, detecting even small declines in affected populations of  
438 animals is of paramount importance. Given the relatively higher quality data produced at lower financial  
439 and human costs, when the hierarchical spatial/temporal survey style is not economically feasible, the  
440 recommended survey method for detecting changes in the occupancy of streams by Arctic Grayling  
441 young-of-year is streamside visual surveys.

442

#### 443 **Lake dewatering effects on downstream populations of Arctic Grayling**

444 Results from the hierarchical model showed that during the summer of 2014, prior to the  
445 dewatering of the upstream Kennady Lake, the probability that streams in the KLM system were  
446 occupied by Arctic Grayling young-of-year was on average 78 - 89% (Fig. 3A) (each data range in this  
447 section gives the estimate from the streamside visual method followed by the estimate from the  
448 electrofishing method from the hierarchical model). There was a 28 -38% chance that a stream would  
449 become unoccupied by fish between survey periods. There was a fairly narrow range of abiotic and  
450 biotic conditions in the KLM streams during the 2014 surveys (see Table S1); conditions were relatively  
451 uniform across streams and were well within the ranges reported by Jones & Tonn (2004) as being  
452 suitable for use by young-of-year Arctic Grayling in Barrenland streams. Water velocity in streams early  
453 in the season (Fig. 5A) had the strongest influence on the probability of occupancy, with decreasing  
454 probability of occupancy as water velocities increased from 0.05 to 0.2 m/s. This is nearly identical to  
455 the findings of a previous study conducted on Arctic Grayling young-of-year (Jones and Tonn 2004).  
456 Overall, the ranges of ideal depths and velocities in streams in the KLM system in 2014 provided a great  
457 deal of suitable Arctic Grayling young-of-year rearing habitat.

458 After the start of the dewatering of Kennady Lake in 2015, the probability of streams being  
459 occupied by Arctic Grayling young-of-year in the affected streams (KLM system) was lower. The

460 probability that streams were occupied was only 31-39% (compared to 78-89% prior to dewatering).  
461 There was negligible probability of occupancy in several streams of the KLM system, including K5 (the  
462 first stream immediately downstream of Kennady Lake), and only very small portions of streams M3, M2  
463 and M1 were likely to have been occupied (Fig. 3B). The probability of occupancy in the downstream  
464 control P system and the unconnected control W system was higher than in the affected KLM system;  
465 averaging 48-82% and 40-73% respectively, despite these streams having otherwise similar habitat  
466 characteristics to the KLM system (see Table S1). Only one control stream, P8 had no observations of  
467 Arctic Grayling young-of-year. Stream P8 was also the deepest (60 cm) and had the fastest average  
468 water velocity (1.02 m/s) of all the streams sampled in 2015. Within the affected KLM system, stream  
469 L1B had the highest probability of occupancy (94%, Fig. 3B) and was also both the shallowest stream in  
470 July (average of 24 cm), and had the lowest amount of connected wetlands (16%).

471           Greater water depths early in the open-water season (early July) likely reduced the probability  
472 that Arctic Grayling young-of-year would occupy a stream throughout the summer of 2015 (Figs 5B and  
473 5C). Unlike the conditions prior to dewatering in 2014, after dewatering activities had begun there was a  
474 slightly positive relationship between increasing water velocity and Arctic Grayling young-of-year  
475 probability of occupancy (Fig. 5C). Water velocities in 2015 were a great deal higher in the affected KLM  
476 system than the previous year (see Table S1), although within the range of water velocities observed in  
477 the two control (P and W) systems. Jones and Tonn (2004) reported that there was clearly a weaker  
478 preference for optimal water velocities ( $0.1 \text{ m s}^{-1}$ ) over optimal water depths (10-20 cm) for larger Arctic  
479 Grayling young-of-year in Barrenlands streams. With shallow habitats in short supply, Arctic Grayling  
480 young-of-year may have been forced to tolerate less optimal water flows in the shallowest areas of the  
481 streams in favour of avoiding predators, such as Northern Pike, inhabiting deeper water.

482           Although water flow in the affected streams in the KLM system in 2015 was within the range  
483 observed in control streams, the area of wetlands surrounding the streams and amount of submerged  
484 vegetation (vegetation may have become submerged as water moved out laterally from the flooded  
485 streams) was higher in the affected KLM system than in the control streams. Where stream water was  
486 deeper, there was a greater prevalence of floodplain wetlands along the sides of the stream channel  
487 ( $R^2=0.50$ , Fig. 5B), coincident with a declining probability that young-of-year Arctic Grayling were present  
488 in the stream. The area of wetlands surrounding the streams in the KLM system was much higher after  
489 lake dewatering began in 2015 (16-50%) than prior to dewatering in 2014 (0-34%). Given the relatively  
490 flat landscape, the excess discharge of water into the system as a result of the dewatering of Kennady  
491 Lake likely moved out laterally from streambanks of the affected KLM streams, to fill wetlands instead of  
492 significantly increasing the depth of these streams. Based on the findings of this study, the persistence  
493 of this Barrenlands population of Arctic Grayling should be ensured if appropriate, site-specific targets of  
494 minimum, maximum and ideal water depths and velocities are established and closely monitored in  
495 streams immediately downstream of dewatering activities.

496

#### 497 **Implications for monitoring**

498           The design of a cost-effective monitoring plan is crucial to the protection of animals in areas  
499 affected by anthropogenic activities. An occupancy-modelling framework was used to provide guidance  
500 on several common issues in the prediction of habitat use by a sentinel and valued species of fish,  
501 including imperfect detection, appropriate sampling methods and the best allocation of efforts spatially  
502 and temporally during very short seasons in difficult and remote terrain. Currently, we know of no other  
503 studies that have addressed issues of sequential spatial correlation in occupancy modelling of stream  
504 fish populations, given the inevitability that the standard method of observing fish while walking along

505 the stream will produce spatially auto-correlated, and thus biased, estimates of site occupancy (Hines et  
506 al. 2010).

507         Streamside visual surveys performed similarly to electrofishing surveys for Arctic Grayling  
508 young-of-year in these shallow streams. Given the minimal potential for injury to imperilled fish  
509 populations, we recommend the use of the less invasive streamside visual survey method over  
510 electrofishing for occupancy models of fish in non-turbid streams. The comparison of spatially-replicated  
511 with temporally-replicated occupancy study designs showed that the spatially-replicated model  
512 produced probabilities of occupancy that were the least biased compared to the (best) hierarchical  
513 model. Most importantly, the spatially-replicated data set was capable of detecting the decline in the  
514 probability of occupancy in the streams affected by mining operations, whereas the temporally-  
515 replicated surveys could not. When there is a great need to survey large expanses of rough, remote  
516 terrain, there is often a trade-off in allocation of effort. Facing a decision between spatially-replicated or  
517 temporally-replicated surveys, we found that spatial replication can provide suitably sensitive, time and  
518 cost-effective standardized data sets for modelling the probability of stream occupancy of Arctic fishes.

519         Surveys of adjacent stream segments produced spatially correlated data, as expected. This issue  
520 was alleviated by surveying randomly-selected stream segments within a hierarchical spatially- and  
521 temporally-replicated occupancy model. The assumption that a site is closed to changes in the  
522 probability of occupancy in spatially-replicated surveys requires that the species' home range is similar  
523 to the size of the site, such that the species is available for detection in all of the spatial replicates within  
524 an occupied site (Kendall and White 2009). Positive bias in occupancy probabilities is commonly  
525 observed in spatial surveys where sites are sampled exhaustively and/or without replacement (Kendall  
526 1999, Kendall and White 2009), leading to the overestimation of occupancy probabilities. This can  
527 increase the chances that a real decline in a population will go undetected. Currently, we know of no

528 other studies that have addressed issues of sequential spatial correlation in occupancy modelling of  
529 stream fish populations, given the near certainty that the standard method of observing fish while  
530 walking along the stream will produce spatially auto-correlated, and thus biased, estimates of site  
531 occupancy (Hines et al. 2010). Iterations of occupancy models have been designed to account for  
532 violation of the assumption of independent observations (spatial-autocorrelation), and to account for  
533 violation of the assumption of closure in time (staggered entry models), but there is no model that  
534 allows for the violation of the assumption of closure in space. We suggest that it may be better to avoid  
535 violating the spatial closure assumption by relaxing the effect of violating the assumption of  
536 independence of spatial replicates. A downstream sequential survey style, but with replicate surveys  
537 taking place immediately downstream of the first observation, might help to better meet the  
538 assumption that fish are available in all segments, when present in the stream. Correlation in the  
539 presence of Arctic Grayling young-of-year in sequential segments likely arises from the poor swimming  
540 ability of a group of newly hatched fry that are easily displaced downstream during early larval stages  
541 (Deleray and Kaya 1992).

542         The observed absence of young-of-year Arctic Grayling in some spatial replicates of otherwise  
543 occupied streams suggests that they may use smaller suitable patches within a stream, as opposed to  
544 occupying the entire length of a stream. If true, mitigating the effects of anthropogenic alterations of  
545 water flow on whole streams may be less important than ensuring that a smaller portion of the stream  
546 be maintained as suitable rearing habitat for Arctic Grayling young-of-year and that other portions of  
547 the stream simply remain passable for fish migration.

548

549 **Acknowledgements**



550 This study was funded by a DeBeers Canada Inc., a Collaborative Research and Development Grant from  
551 the Natural Sciences and Engineering Research Council of Canada (to HKS), a Canadian Graduate  
552 Scholarship from the Natural Sciences and Engineering Research Council of Canada (to K.J.A.), and a  
553 Northern Scientific Training Program grant (to K.J.A.). The field program was executed with assistance  
554 from S Lord, M Elwood, C Barrett, C Davis, D Swift, J Fregoe, S Catholique, J Remshardt, D Rusch, S  
555 McLeod, C Marsh, and K Allen. M Servos and M Power provided helpful comments on an early draft of  
556 this manuscript, and N Mochnacz and D Mackenzie aided with early study design.

557

#### 558 **Data accessibility**

559 Data used in this study will be accessible from the University of Waterloo's institutional repository  
560 searchable from: <https://uwspace.uwaterloo.ca/handle/10012/9937> (data will be deposited at time of  
561 article acceptance with a direct address provided here).

562

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661

662 **Figure Captions**

663 **Fig. 1.** Location of study area.

664

665 **Fig. 2.** Comparison of the detection probabilities  $\pm$  95% CI estimated from two common fish sampling  
666 methods in A) 2014 and B) 2015. Detection probabilities were affected by C) water velocity (m/s) in  
667 2014, and D) water depth (m) in 2015.

668

669 **Fig. 3.** Probability of occupancy  $\pm$  SE vs. observed presence of Arctic Grayling young-of-year in streams in  
670 A) 2014 and B) 2015 based on survey method.

671

672 **Fig. 4.** Relative bias in probabilities  $\pm$  95% CI of A) detection and B) occupancy produced by simulated  
673 occupancy surveys of only spatial or only temporal replicates compared to the combined hierarchical  
674 model including both spatial and temporal replication.

675

676 **Fig. 5.** Proportion of streams occupied modelled as a function of influential habitat variables based on  
677 data from A) 2014 streamside visual surveys, B) 2015 streamside visual surveys, and C) 2015  
678 electrofishing surveys. No habitat variables were found to significantly influence the probability of  
679 occupancy of streams in 2014 electrofishing surveys.

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683 **Supporting information**

684 Additional Supporting Information may be found in the online version of this article:

685 Supplemental Data S1. Description and rationale for use of *a priori* candidate hierarchical occupancy  
686 models.

687 Table S1. Ranges of sampled habitat variables.

688 Table S2. Beta coefficients for benchmark hierarchical models.

689 Fig. S1. Conceptual representation of the data configurations used of the analysis of presence-absence.

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699 Table 1. Benchmark model selection results of *a priori* candidate models of the 2014-visual hierarchical,  
700 2015-electrofishing hierarchical, 2015-visual hierarchical and, 2015-electrofishing hierarchical data sets  
701 of presence-absence of Arctic Grayling young-of-year in Barrenland streams.  $\Delta AICc$  (corrected for small  
702 sample size) of each model from the minimum model was used to calculate AICc weight ( $w_i$ ) and rank of  
703 each model, and to select the most parsimonious model from the three candidate structures.  $K$  is the  
704 number of model parameters, and  $-2\log(L)$  is the negative of twice the logarithm of the likelihood

705 function evaluated at the maximum likelihood estimates. \* indicates models that failed to converge  
706 mathematically and were removed from consideration